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Gender health care inequalities in health crisis: when uncertainty can lead to inequality



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Abstract

Background In health crisis, inequalities in access to and use of health care services become more evident. The objective of this study is to analyse the existence and evolution of gender inequalities in access to and use of health-care services in the context of the COVID-19 health crisis.

Methods Retrospective cohort study using data from all individuals with a confirmed COVID-19 infection from March 2020 to March 2022 in Aragón (Spain) (390,099 cases). Health care access and use was analysed by gender for the different pandemic waves. Univariate and multivariate analyses were conducted to evaluate the effect of sex in health care. Blinder-Oaxaca decomposition methods were performed to explain gender gaps observed.

Results The health care received throughout the COVID-19 pandemic differed between men and women. Women were admitted to hospital and intensive care units less frequently than men and their stays were shorter. Differences observed between men and women narrowed throughout the pandemic, but persisted even after adjusting for age, socioeconomic status, morbidity burden or the patient's place of residence. Differences in sociodemographic characteristics and morbidity burden could explain partially the gender inequalities found, mainly in the later phases of the pandemic, but not in the earlier waves.

Conclusions There were gender inequalities in access to and use of health services during the COVID-19 pandemic. Inequalities were greater in the first waves of the pandemic, but did not disappear. Analysis of health crises must take into account an intersectional gender perspective to ensure equitable health care.

Keywords Gender inequalities, Health crisis, COVID-19, Health care use

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Text box 1. Contributions to the literature

• Despite the importance of the gender approach to health crisis management, which involves considering gender inequalities in policy analysis, planning, design and implementation, it is not sufficiently applied

• We observed gender inequalities in the access to and use of health services during the COVID-19 pandemic. These inequalities were greater in the first waves of the pandemic, but did not disappear

 Preparing for future pandemics and health crises by focusing on care for the most vulnerable groups is necessary to avoid exacerbating existing inequalities

Background

During health crisis there is a complex relationship between competing pathologies and the social and structural conditions in which social health inequalities are propagated and reinforced [1, 2]. In relation to gender, it is also during crises that large differences in mortality and vulnerability to disease become apparent [3]. But these inequalities are not limited to the risk of contracting the disease, but also to inequalities in access to and use of health care services [4].

The state of alarm generated during the COVID-19 pandemic, as has been described in other pandemics such as Zika or Ebola [5–7], had a differential effect on women. The likelihood of contracting pathologies increased, since women provided most of the informal care within families. As a result, they also experienced a higher frequency of work limitations and reduced income. All these situations mean a greater burden of illness for women for gender reasons, as they are exposed to a greater situation of vulnerability, which entails a differential vulnerability between men and women. This also limits their degree of freedom in decision-making and conditions the practices related to the pandemic that women and men have in their living environment, including their relationship with health care [8].

Although this is a well-known fact, some authors point out that during crisis health data are often incomplete and information is not always disaggregated by sex [9]. This makes it difficult to establish a relationship between sex and susceptibility to the disease, as well as in the health care received. Also, the impact on gender is not sufficiently taken into account, nor are health policy efforts and measures adopted from a public health perspective [10]. In this sense, the United Nations published a report [11] highlighting the importance and necessity of a gender approach in crisis management, which implies considering gender inequalities in the analysis, planning, design and implementation of policies, along with actions to address the different situations and needs that may arise.

To analyse the extent to which health emergencies affect women and men unequally is critical to develop effective and equitable policies and interventions in future crises. In this context, the aim of our study is to analyse the existence and evolution of gender inequalities in access to and use of healthcare services in the context of a health crisis, using the COVID-19 pandemic as a case study.

Methodology

Design, information sources and study population

We conducted a retrospective cohort study using data from the Aragón-COVID19 cohort. This is a health data collection of all individuals undergoing COVID-19 testing in the Spanish region of Aragón since March 2020. Aragón is an Autonomous Community located in the northeast of Spain. It is the fourth Spanish Community by extension but occupies the 11th place of 17 in terms of population. It has a population of 1.3 million inhabitants and half of the population live in the city of Zaragoza. Spain has a public health system which covers practically the entire population.

People included in the Aragón-COVID19 cohort were tested by the public health system either when they had symptoms compatible with COVID-19 or when they had close contact with a confirmed subject. All COVID-19 cases were confirmed using polymerase chain reaction (PCR) or COVID antigen testing. The Aragón-COVID19 cohort includes information gathered from administrative health data sources as well as electronic health records of the Aragón Healthcare Service.

All individuals in the cohort included in this study were those tested from 9 March 2020, the first epidemiological week with COVID-19 cases reported in Aragón, to 31 March 2022, which corresponded to the seventh wave in Aragón.

Variables of the study

Of all the individuals in the cohort we analyzed their sociodemographic and clinical information, and their use of healthcare services related with the COVID-19 episode.

The main variable analysed was sex, obtained from the Users Database of Aragon Healthcare Service. Regarding sociodemographic characteristics, we also considered the age (under 15, 15–44, 45–64, 65–79 and 80 years or older), socioeconomic level and information about the place of residence. Socioeconomic level was calculated on the basis of pharmacy copayment levels and social security benefits received. From the combination of these two variables, 7 mutually exclusive categories were obtained: employed individuals earning less than €18,000 per year, employed individuals earning €18,000 per year or more, individuals receiving the unemployment allowance, individuals with a contributory pension of less than €18,000

per year and receiving free medicines, individuals with a contributory pension of €18,000 per year or more, individuals affiliated to the mutual insurance system for civil servants and other situations not previously considered. We considered also some variables related with the place of residence of the patient. Residing in a long-term care (LTC) facility was considered. We described deprivation index of the Basic Healthcare Area of residence categorized into four quartiles, from least (Q1) to most (Q4) deprived [12]. Finally, we considered if patient resided in a rural or urban area, according to the Aragon Government [13].

The clinical information included was obtained from the morbidity adjusted groups (GMA) [14]. This source of information considers all medical diagnoses available in primary healthcare and hospital discharge records (Minimum Basic Data Set of Hospital Discharges). We considered GMA information from January 2020 in order to know the health status prior to the COVID-19 diagnosis of the cohort individuals. The three variables analyzed from GMA were morbidity burden (obtained from the aggregation of the patient's different diagnoses), the presence of chronic morbidities and the presence of respiratory illnesses.

We used different indicators to evaluate healthcare delivery. Firstly, we considered the time from the beginning of COVID-19 symptoms to diagnosis, with a range from -7 to 15 days, in those patients with symptoms. We also calculated the time from the COVID-19 diagnosis to hospital admission (ranging from -15 to 30 days). We analyzed hospital admission (yes/no) and intensive care units (ICU) admission (yes/no) in this hospitalization. As we do not have the diagnosis related to hospitalization, we considered a hospitalization to be related to COVID-19 when it happened between -15 and 30 days of COVID-19 diagnosis. We also calculated the length of hospital stay and the length of ICU stay.

Analyses

All the analyses were stratified by pandemic waves in three groups: wave 1 (from March to 20 June 2020), waves 2 and 3 (from 21 June to 27 December 2020) and waves 4 to 7 (from 28 December 2020 to March 2022). We separated them according to the knowledge of COVID-19 management in each stage, as well as with the epidemiological situation. So, wave 1 is analysed separately due to the high level of uncertainty with diagnosis and clinical management. In waves 2 and 3 the management of disease was clearer, but there was no access to vaccines. Finally, in waves 4 to 7 there was already a COVID-19 vaccine available and the management of disease had improved.

First, we described sociodemographic and clinical characteristics of patients included in the study. We performed descriptive analyses of each health care indicator stratified by sex and by each of the stages of the COVID-19 pandemic, in order to identify differences by gender. To analyse the effect of sex on health care utilization, we performed univariate and multivariate analyses separately by each COVID-19 pandemic stage. To evaluate the association between sex and time from symptoms to confirmation and time from diagnostic to confirmation, linear regression analyses were performed. We conducted logistic regression models to study the influence of sex in hospital and ICU admissions. For the length of hospital stay and length of ICU stay, Poisson regression models were performed. All the analyses were adjusted by age, the presence of comorbidities, socioeconomic level and residence in a long-term care facility, in order to control for these variables.

Sex differences in health care attention were decomposed using Blinder-Oaxaca decomposition method. This method divides the mean outcome observed differences between men and women into two components: the explained component, that captures the differences in the outcomes explained by the independent variables, and the unexplained component, that also captures all potential effects of unobserved variables. [Oaxaca RL: Male– female wage differentials in urban labor markets. Int Econ Rev. 1973, 14 (3): 693–709. https://doi.org/10.2307/ 2525981.]. A twofold decomposition applying Oaxaca R library and reference regression coefficients were calculated from a pooled regression model [https://cran.rproject.org/web/packages/oaxaca/vignettes/oaxaca.pdf].

Analyses were conducted using R 4.1.3 (2022–03-10). All data were pseudonymized and the research protocol was approved by The Clinical Research Ethics Committee of Aragón (CEICA) (PI20/184).

Results

From March 2020 to March 2022 390,099 confirmed positive cases of COVID-19 were diagnosed in Aragón. 52.3% of the cases were diagnosed in women.

The most frequent group of age was those from 16 to 44 years old. In the older age group (≥ 80 years), women accounted for a higher percentage. By socioeconomic level, the highest percentage in men was in those employed with salaries higher than 18,000€ per year (34.4%). In women, it belonged to those women employed earning less than 18,000€ per year (38.5%). According to the place of residence, more women lived in LTC facilities and a urban predominance was observed. Related to health status, women had a higher morbidity burden than men (p < 0.001). There were differences by sexes for all the diseases studied, with the only exception of chronic kidney disease. Men had a higher prevalence than women of diabetes, ischemic heart disease, stroke, hypertension, chronic obstructive pulmonary disease (COPD), cirrhosis and HIV/AIDS. On the contrary, women showed a higher prevalence of heart failure, depression and dementia (Table 1).

We analysed the evolution across the COVID-19 pandemic of different health care indicators by sex. Time from symptoms to diagnosis was higher in men than in women during the first wave (mean: 4.7 days vs. 3.9; p < 0.001), but differences disappeared in following waves. On the contrary, time from diagnosis to hospital admission was lower in men than in women, and these differences remained for all the period (p = 0.001). The frequency of hospital admission and ICU admission was higher in men than in women especially in the first wave (53.1% vs 35.6% for hospital admission and 6.8% vs. 2.2% for ICU admission). These differences decreased in the following waves but persisted for all the pandemic (p < 0.001). Finally, length of hospital and ICU stay was

Table 1 Sociodemographic and clinical characteristics of COVID-19 confirmed positive cases by sex (march 2020-march 2022) inAragón, Spain

| | Men (N=186,455) | Women (N = 203,644) | p |
|--|-----------------|---------------------|---------|
| Age (years) ^a | | | < 0.001 |
| < = 15 | 34,028 (18.2%) | 32,573 (16.0%) | |
| 16–44 | 74,331 (39.9%) | 82,414 (40.5%) | |
| 45–64 | 52,002 (27.9%) | 55,854 (27.4%) | |
| 65–79 | 17,022 (9.13%) | 17,801 (8.74%) | |
| ≥80 | 9072 (4.87%) | 15,002 (7.37%) | |
| Socioeconomic level ^a | | | < 0.001 |
| Employed < 18,000€ per year | 62,249 (33.4%) | 78,324 (38.5%) | |
| Employed ≥ 18,000€ per year | 64,099 (34.4%) | 49,319 (24.2%) | |
| Unemployed | 5458 (2.93%) | 8521 (4.18%) | |
| Pensioner < 18,000€ per year | 23,736 (12.7%) | 34,979 (17.2%) | |
| Pensioner ≥ 18,000€ per year | 12,597 (6.76%) | 9507 (4.67%) | |
| Mutualist | 8319 (4.46%) | 8200 (4.03%) | |
| Other | 9997 (5.36%) | 14,794 (7.26%) | |
| Residing in a long-term care facility ^a | 4237 (2.27%) | 7707 (3.78%) | < 0.001 |
| Deprivation quartile ^a | | | 0.031 |
| 1 (least deprivation) | 52,969 (28.8%) | 58,772 (29.1%) | |
| 2 | 44,674 (24.3%) | 49,222 (24.4%) | |
| 3 | 37,291 (20.3%) | 40,264 (20.0%) | |
| 4 (highest deprivation) | 48,779 (26.6%) | 53,469 (26.5%) | |
| Zone of residence ^a | | | < 0.001 |
| Rural | 52,403 (28.5%) | 54,131 (26.8%) | |
| Urban | 131,310 (71.5%) | 147,596 (73.2%) | |
| Morbidity burden ^b | 3.36 (3.97) | 4.01 (4.19) | < 0.001 |
| Morbidity ^a | | | |
| Diabetes Mellitus | 11,465 (6.49%) | 9421 (4.83%) | < 0.001 |
| Heart failure | 1891 (1.07%) | 2553 (1.31%) | < 0.001 |
| Ischemic heart disease | 4905 (2.78%) | 2362 (1.21%) | < 0.001 |
| Stroke | 2415 (1.37%) | 2312 (1.19%) | < 0.001 |
| Hypertension | 27,430 (15.5%) | 28,907 (14.8%) | < 0.001 |
| COPD | 4557 (2.58%) | 2565 (1.31%) | < 0.001 |
| Chronic kidney disease | 6212 (3.51%) | 6980 (3.58%) | 0.299 |
| Cirrhosis | 2754 (1.56%) | 1985 (1.02%) | < 0.001 |
| HIV/AIDS | 429 (0.24%) | 252 (0.13%) | < 0.001 |
| Depression | 8233 (4.66%) | 21,944 (11.2%) | < 0.001 |
| Dementia | 1725 (0.98%) | 4246 (2.18%) | < 0.001 |

N number, p statistical significance.^aNumber (percentage).^bMean (Standard Deviation). COPD chronic obstructive pulmonary disease

higher in men than in women, and these differences persisted for the whole period analysed (Fig. 1).

We conducted univariate and multivariate models in order to know the effect of sex in the different health care indicators evaluated, adjusting by age, morbidity burden, socioeconomic level and residence in a long-term care facility (Table 2).

The time from symptom onset to diagnosis was shorter in women than in men in waves 1 and 2–3. Specifically, COVID-19 diagnosis in women showed a reduction of 0.82 days compared to men in wave 1. These differences decreased throughout the pandemic, showing no statistically significant differences in waves 2 to 7 (OR: 0.01, 95% CI: -0.01–0.04 in later waves). In contrast, the time from diagnosis to hospital admission was longer in women than in men. Thus, in women, the greatest differences were shown in the first wave, when they took 0.57 days longer to be admitted to hospital than men (95% CI: 0.28 to 0.87 days). These differences remained practically unchanged during the whole period analysed (OR: 0.52; 95% CI: 0.31 to 0.72 in the last waves).

The risk of hospital and ICU admission was lower in women than in men and remained statistically significant for the whole period, with higher differences in the risk of ICU admission (OR: 0.41; 95%CI: 0.35–0.48 in the last waves). Regarding the length of stay, it was lower in women than in men. In the case of length of hospital stay differences remained stable for all the period analysed (OR: 0.86; 95%CI 0.86–0.87 in the last waves). On the contrary, differences between men and women regarding length of ICU stay decreased across the pandemic, but remained statistically significant (OR: 0.82 in the first wave vs. 0.93 in the last waves).

Time to diagnosis and time to hospital admission: linear regression model; Hospital admission and ICU admission: logistic regression model; Length of hospital stay (days) and length ICU stay (days): Poisson model; Analyses adjusted by age, morbidity burden, socioeconomic level, area of residence (urban or rural) and residence in a long-term care facility.

In order to explain the gender gaps observed, Oaxaca decomposition analyses were conducted for each wave (Table 3). Variables included in the analyses were age, morbidity burden, socioeconomic level, area of residence (urban or rural) and residence in a LTC facility. Explained fraction of the gender gap increased across the pandemic for all the health care indicators analysed, with the only exception of length of hospital stay and length of ICU stay. The health care indicator with the highest explained estimate was time to diagnosis at waves 4–7 (52.8%) followed by hospital admission (31.11%).

Regarding the contribution of each variable, age played an important role in gender health care differences for hospital and ICU admission. Socioeconomic level explained also a high portion of the gender health care gap in hospital admission and ICU admission, especially in waves 4-7 (30.1% in hospital admission and 43.9% in ICU admission). It was also an important factor explaining gender gaps in the length of ICU stay in wave 1 (56.1%) and waves 2-3 (41.6%) and for the time from diagnosis to hospital admission across all the waves. Residing in a LTC facility explained a high proportion of gender health care differences for ICU admission, time to diagnosis, time to hospital admission and length of hospital stay, especially during the first wave. Finally, differences in morbidity burden were only important to explain gender differences in the length of hospital stay across all the period analysed (Supplemental Figs. 1, 2 and 3).

Discussion

Men and women with COVID-19 infection had different profiles: women were generally older than men, had lower socioeconomic status, lived in LTC facilities more frequently and had a higher morbidity burden. Women also had a higher frequency of COVID-19 infection than men.

The health care received throughout the pandemic differed between men and women, with gender inequalities being observed. Women were diagnosed earlier than men, but these differences disappear in the last waves of the COVID-19 pandemic. Women were admitted to hospital and ICU less frequently than men and their stays were also shorter. Overall, these differences narrowed throughout the pandemic, but persisted for all the period analysed, even after adjusting for other characteristics such as age, socioeconomic status, morbidity burden or the patient's place of residence. Differences in sociodemographic characteristics and morbidity between women and men could explain partially the gender differences found, mainly in the later phases of the pandemic, but not in the first waves. Particularly striking is the importance of residing in an institution in the early phases and the high importance of socioeconomic status throughout the pandemic as explanatory factors for the gender inequalities found.

As widely described, the incidence of COVID-19 in women was higher than in men [15]. There are a number of circumstances that have been associated with these gender differences, such as genetic or hormonal factors [16], but also other factors that generate greater vulnerability in women, such as their lower socio-economic status [17].

Gender inequalities in health care have been observed. Men are admitted more often to hospital and ICU and



Fig. 1 Evolution of health care indicators by sex and wave during the COVID-19 pandemic (march 2020-march 2022) in Aragón, Spain

Table 2 Risk of women in relation to men for health care delivery by pandemic wave during the COVID-19 pandemic (march 2020march 2022) in Aragón, Spain. Univariate and multivariate models

| | WAVE 1 | | | | WAVE 2–3 | | | | WAVE 4–7 | | | |
|-----------------------------------|---------------------|-----------|-----------------------|-----------|------------------|-------------|-----------------------|------------|---------------------|-----------|-----------------------|------------|
| | Univariate model | | Multivariate model | | Univariate model | | Multivariate model | | Univariate model | | Multivariate model | |
| | OR | 95%Cl | OR | 95%CI | OR | 95%Cl | OR | 95%Cl | OR | 95%Cl | OR | 95%Cl |
| Time to diagnosis (days) | -0.82 | -1.090.56 | -0.63 | -0.890.37 | -0.06 | -0.12-0.000 | -0.05 | -0.11-0.01 | 0.03 | 0.00-0.05 | 0.01 | -0.01-0.04 |
| Time to hospital admission (days) | 0.57 | 0.28-0.87 | 0.53 | 0.23-0.84 | 0.36 | 0.13-0.60 | 0.55 | 0.30-0.79 | 0.32 | 0.13-0.52 | 0.52 | 0.31-0.72 |
| Hospital admission | 0.49 | 0.44-0.55 | 0.44 | 0.38-0.50 | 0.76 | 0.72-0.80 | 0.57 | 0.53-0.60 | 0.79 | 0.76-0.82 | 0.62 | 0.60-0.65 |
| Lenght of hospital stay (days) | 0.89 | 0.87-0.91 | 0.86 | 0.84-0.88 | 0.89 | 0.88-0.90 | 0.88 | 0.87-0.89 | 0.88 | 0.87-0.89 | 0.86 | 0.86-0.87 |
| ICU admission | 0.3 | 0.22-0.40 | 0.37 | 0.26-0.51 | 0.42 | 0.35-0.51 | 0.42 | 0.34-0.50 | 0.47 | 0.40-0.54 | 0.41 | 0.35-0.48 |
| Lenght of ICU stay (days) | 0.87 | 0.81-0.93 | 0.82 | 0.76-0.88 | 0.93 | 0.89–0.97 | 0.95 | 0.91-0.99 | 0.97 | 0.94-1.00 | 0.93 | 0.90-0.96 |

Table 3 Decomposition of gender inequality in health care delivery by pandemic waves during the COVID-19 pandemic (march 2020-march 2022) in Aragón, Spain. Oaxaca decomposition analyses

| | Wave 1 | | Waves 2–3 | | Waves 4–7 | | |
|---------------------|----------|--------------|-----------|--------------|-----------|--------------|--|
| | Absolute | Relative (%) | Absolute | Relative (%) | Absolute | Relative (%) | |
| Time to diagnosis | | | | | | | |
| Explained | 0.006 | 0.739 | 0.008 | 16.390 | -0.015 | 52.789 | |
| Unexplained | 0.794 | 99.261 | 0.043 | 83.610 | -0.0137 | 47.211 | |
| Time to hospital ad | mission | | | | | | |
| Explained | 0.004 | 0.699 | 0.135 | 21.072 | 0.145 | 24.266 | |
| Unexplained | -0.573 | 99.301 | -0.507 | 78.928 | -0.451 | 75.734 | |
| Hospital admission | l | | | | | | |
| Explained | 0.093 | 8.950 | -0.174 | 25.955 | -0.159 | 31.114 | |
| Unexplained | 0.942 | 91.050 | 0.495 | 74.045 | 0.351 | 68.886 | |
| Lenght of hospital | stay | | | | | | |
| Explained | -0.938 | 23.840 | -1.135 | 29.770 | -0.507 | 17.756 | |
| Unexplained | 2997 | 76.160 | 2678 | 70.230 | 2347 | 82.244 | |
| ICU admission | | | | | | | |
| Explained | 0.087 | 5.879 | -0.070 | 7.108 | -0.141 | 14.318 | |
| Unexplained | 1.393 | 94.121 | 0.916 | 92.892 | 0.846 | 85.682 | |
| Lenght of ICU stay | | | | | | | |
| Explained | -0.636 | 12.582 | 0.480 | 28.910 | -0.253 | 17.523 | |
| Unexplained | 4.418 | 87.418 | 1.180 | 71.090 | 1.190 | 82.477 | |
| | | | | | | | |

Variables included in the Oaxaca decomposition analyses: age, morbidity burden, socioeconomic level, area of residence (urban or rural) and residence in a long-term care facility

stay longer in hospital, which results in a greater therapeutic effort. This phenomenon is well known and, unfortunately, is not limited to health crises. It has already been observed that, facing the same health problem, hospitalization rates are generally lower in women than in men [18, 19] which suggest that women could face more obstacles in accessing health care. So, women tend to stay at home, rather than being hospitalized and the length of hospital stay is generally shorter [20], which could be explained by their caring role, as it has been already described [21]. But there are other factors that may be involved in gender inequalities in health care, like health attitudes, health behaviours and health care needs [22]. In the context of the COVID-19 pandemic there are also other circumstances that may have contributed to gender inequalities. Men are usually cared by women in their homes but women, often older, are referred to nursing homes when they are ill. Women living in LTC facilities are in a situation of greater vulnerability, which increases if they also suffer from dementia, which has been associated in this context with a higher risk of not being admitted to hospital [23]. The greatest impact of the COVID-19 pandemic on institutionalized people has been associated with physical and psychological factors, living conditions and deficient policy responses [24–26], which seems to be associated with difficulty in accessing health care. Finally, women have lower socioeconomic status than men, a factor classically associated with poorer health care attention [27, 28].

Differences in all indicators of access to and use of health services are greater at the beginning of the pandemic and tend to decrease as the pandemic progresses, but do not disappear completely. These results show that, in times of health system crisis and great uncertainty, existing inequalities are exacerbated. In the case of COVID-19, the lack of clear protocols for action and a limited definition of the disease in the early stages of the pandemic that did not take gender differences into account may explain these greater inequalities in utilisation [29]. Moreover, as some authors point out [29], hospital and ICU admission are indicators of the severity of the disease, but also of the diagnostic and therapeutic effort required. In our study, and as it has been described in the literature [30], men had higher mortality rates than women. However, when we described the use of health services only in patients who died within 30 days of diagnosis, the gender inequalities in the use of services remained (results not shown). In these sub-analyses, men also showed a higher frequency of hospital admission than women, for all three time points analysed. Thus, in wave 1, men who died within 30 days of COVID-19 diagnosis had a hospital admission in 76.7% of cases, while in women this percentage was 69.6% (p = 0.031). The length of hospital stay was also longer in men, with statistically significant differences for the three time points analysed. Finally, admission to the ICU was also more frequent in men than in women, with the greatest differences in wave 1 (10.7% in men vs. 3.6% in women). This fact has been already observed in Spain, where the probability of admission to the ICU was higher in deceased men than in deceased women, which could indicate, among others, a greater therapeutic effort in men than in women, implying a gender bias in health care [31].

In addition, it should not be overlooked that some authors point to a higher percentage of undiagnosed excess COVID-19 mortality in women, related with an underreporting of COVID-19 deaths among women who died in LTC facilities, especially during the first wave [32], a higher frequency of deaths in women from other causes consistent with COVID-19 [31] and a greater hospital access and care in men [33] that have implied higher reporting of deaths in men in the pandemic. This aspect requires further attention and underlines the need for a gender-sensitive definition of the disease and for sex-disaggregated information in future pandemics [34].

In the first pandemic waves, especially in the first wave, when gender inequalities are greater, these inequalities are not justified either by sociodemographic factors or by morbidity factors of the subjects. As mentioned above, the confluence of factors such as older age, greater morbidity burden, lower socio-economic status and residence in a LTC facility explain, at least partly, the differences found between men and women in the later phases of the pandemic, but not at the beginning of the pandemic. Several factors could explain these differences between the early and late phases of the pandemic. Confinement may have made it particularly difficult for women to seek medical care, due to the burden it entails, which may have prevented them from seeking medical assistance [35]. There were also situations of gender-based violence, in which women did not seek medical care [8, 36]. Finally, the great heterogeneity of existing protocols in the initial phases may have increased the existing gender bias [4].

This study has some limitations. The sex disaggregation available in the electronic registers is binary in nature, which does not necessarily equate to the sex of the subject, and does not allow for the identification of persons with diverse gender identities. Unfortunately, this is the only information available. There are also some limitations inherent to observational studies, such as the existence of incomplete information. In this sense, quality of the data may have changed across the waves. The cause of hospital admission was not available. In order to solve this problem, a range from -15 to 30 days from hospital admission to COVID-19 diagnosis was applied. Some of the patients living in LTC facilities who were not hospitalized could have been treated in one of the "COVID centers" set up in Aragón in the first waves of the pandemic. Unfortunately, this information is not available. Other information not available was the severity of the infection. This aspect is relevant as it could have conditioned the use of health services by patients, mainly aspects such as the decision to be hospitalized or to be admitted to the ICU. However, as noted above, we replicated our analyses only in patients who died within 30 days of diagnosis, as a proxy for high-severity COVID-19 cases, with results similar to those in the general population. Finally, it is important to note that what is statistically significant is not always clinically relevant, and further studies are needed to understand the medical implications of the observed inequalities in health care utilization. On the other hand, this work has many strengths. This is a populational study based on a European population of 1.3 million people. We used data from administrative health data sources combined with electronic health records.

Conclusions

Access to and use of health services during the COVID-19 pandemic has shown gender inequalities. These inequalities were greater in the first waves of the pandemic, but did not disappear. Differences between men and women could not be explained in the first waves by socio-demographic factors or by the morbidity burden of the patients, showing a different therapeutic effort. Moreover, the lack of clear protocols in a context of great uncertainty and health crisis seems detrimental to the equitable use of health services and may lead to less therapeutic effort in groups in vulnerable circumstances.

It is the task of public administrations to promote equity in accessibility and in therapeutic effort, in all aspects, including care, regardless of gender, in order to avoid inequalities in health. These differences are striking, as there is a legal regulation in Spain to avoid them [37]. On the other hand, it should be noted that inequalities in health are mainly due to the social determinants of health [38]. And although the main axes of inequality seem to be at the same level and to operate independently, it is necessary to study how the different opportunities of women and men interact with their socio-economic position, educational level or origin from an intersectional perspective.

Therefore, in order to reduce health inequalities between women and men, health research must take into account an intersectional gender perspective, which considers all risk factors and discriminatory factors that put women in a vulnerable situation [39]. Clear gender-sensitive guidelines, gender-sensitive definitions of disease, or provide appropriately disaggregated information, are essential to ensure equitable and quality health care. In addition, the inclusion of a gender perspective in decision-making processes is crucial for effective crisis response and recovery [40]. In this regard, efforts must be made to prepare for future pandemics and health crises by focusing on care for the most vulnerable groups, to ensure that existing inequalities are not exacerbated.

Abbreviations

| COPD | Chronic obstructive pulmonary disease |
|-------|---------------------------------------|
| LTC | Long-term care |
| ICU | Intensive care unit |
| OR | Odds ratio |
| 95%CI | 95% Confidence interval |
| | |

Supplementary Information

The online version contains supplementary material available at https://doi. org/10.1186/s13690-024-01276-7.

Supplementary Material 1.

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Authors' contributions

Conception and design of the study was made by I.A-P and MJ.R. Analyses were conducted by I.A-P, S.C-F and J.T. The draft was written by I.A-P and B.O-A and revised by S.M. All the authors approved the submitted version.

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Availability of data and materials

No datasets were generated or analysed during the current study.

Declarations

Ethics approval and consent to participate

All data were pseudonymized and the research protocol was approved by The Clinical Research Ethics Committee of Aragón (CEICA) (PI20/184).

Consent for publication

Not applicable.

Competing interests

The authors declare no competing interests.

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References

- 1. Horton R. Offline: COVID-19 is not a pandemic. Lancet. 2020;396(10255):874.
- Bambra C, Smith K, Lynch J. The unequal pandemic: COVID-19 and health inequalities. Bristol: Policy Press; 2021.
- Team TNCPERE. The Epidemiological Characteristics of an Outbreak of 2019 Novel Coronavirus Diseases (COVID-19) — China, 2020. China CDC Wkly. 2020;2(8):113.
- Bautista-González E, Werner-Sunderland J, Pérez-Duarte Mendiola P, Esquinca-Enríquez-de-la-Fuente CJ, Bautista-Reyes D, Maciel-Gutiérrez MF, et al. Health-care guidelines and policies during the COVID-19 pandemic in Mexico: A case of health-inequalities. Heal policy Open. 2021;2:100025.
- 5. Boniol M, et al. Gender equity in the health workforce: analysis of 104 countries. Geneva: World Health Organization; 2019.
- Davies SE, Bennett B. A gendered human rights analysis of Ebola and Zika: Locating gender in global health emergencies. International Affairs. 2016;92:1041–60.
- 7. Harman S. Ebola, gender and conspicuously invisible women in global health governance. Third World Q. 2016;37(3):524–41.
- Enguita-Fernàndez C, Marbán-Castro E, Manders O, Maxwell L, Matta GC. The COVID-19 epidemic through a gender lens: what if a gender approach had been applied to inform public health measures to fight the COVID-19 pandemic? Soc Anthropol. 2020;28(2):263.
- The COVID-19 Sex-Disaggregated Data Tracker | Global Health 50/50. 2021. Available from: https://globalhealth5050.org/the-sex-gender-andcovid-19-project/the-data-tracker/?explore=variable&variable=Confi rmed+cases
- Smith J. Overcoming the 'tyranny of the urgent': integrating gender into disease outbreak preparedness and response. Gend Dev. 2019;27(2):355–69.
- Crises UH-LP on the GR to H. Protecting humanity from future health crises: report of the High-Level Panel on the Global Response to Health Crises. UN. 2016. Available from: https://digitallibrary.un.org/record/822489
- Compés Dea ML, Olivan Bellido E, Feja Solana C, Aguilar Palacio I, Hombrebueno García-Carpintero Romero Del, G, Adiego Sancho B. Construcción de un índice de privación por zona básica de salud en Aragón a partir de datos de censo de 2011. Rev Esp Salud Publica. 2018;92:e201812087.

- Gobierno de Aragón. Dirección General de Farmacia del Gobierno de Aragón. 2020. Available from: https://www.aragon.es/organismos/depar tamento-de-sanidad/servicio-aragones-de-salud/direccion-de-coord inacion-asistencial/servicio-de-farmacia
- Monterde D, Vela E, Clèries M. Los grupos de morbilidad ajustados: nuevo agrupador de morbilidad poblacional de utilidad en el ámbito de la atención primaria. Atención Primaria. 2016;48(10):674–82.
- 15. Ministerio de Sanidad. Gobierno de España. Situación actual de COVID-19 en España. Available from: https://www.sanidad.gob.es/areas/alertasEme rgenciasSanitarias/alertasActuales/nCov/situacionActual.htm
- Paschou SA, Psaltopoulou T, Halvatsiotis P, Raptis A, Vlachopoulos CV, Dimopoulos MA. Gender differences in COVID-19. Maturitas. 2022;161:72.
- Aguilar-Palacio I, Maldonado L, Malo S, Sánchez-Recio R, Marcos-Campos I, Magallón-Botaya R, et al. COVID-19 Inequalities: Individual and Area Socioeconomic Factors (Aragón, Spain). Int J Environ Res Public Heal. 2021;18(12):6607.
- De Waure C, Bruno S, Furia G, Di Sciullo L, Carovillano S, Specchia ML, et al. Health inequalities: An analysis of hospitalizations with respect to migrant status, gender and geographical area. BMC Int Health Hum Rights. 2014;14(1):1–10.
- Sánchez-Recio R, Alonso JP, Gil-Lacruz A, Aguilar-Palacio I. Inequities in hospitalisation in a South European country: Lessons learned from the last European recession. Health Soc Care Community. 2022;30(4):e1170–82.
- EuroStat. Eurostat online statistical database Hospital discharges and length of stay statistics. European Comission. 2020. Available from: https://ec.europa.eu/eurostat/statistics-explained/index.php?title=Hospital_discharges_and_length_of_stay_statistics&oldid=585449#Hospital_ discharges_by_sex_and_age
- 21. del Río Lozano M, del Mar García M. Caregiving and the COVID-19 pandemic from a gender perspective. Gac Sanit. 2021;35(6):594–7.
- Gijsbers Van Wijk CMT, Van Vliet KP, Kolk AM. Gender perspectives and quality of care: Towards appropriate and adequate health care for women. Soc Sci Med. 1996;43(5):707–20.
- Aguilar-Palacio I, Maldonado L, Marcos-Campos I, Castel-Feced S, Malo S, Aibar C, et al. Understanding the COVID-19 Pandemic in Nursing Homes (Aragón, Spain): Sociodemographic and Clinical Factors Associated With Hospitalization and Mortality. Front Public Heal. 2022;10:928174.
- Lai C-C, Wang J-H, Ko W-C, Yen M-Y, Lu M-C, Lee C-M, et al. COVID-19 in long-term care facilities: An upcoming threat that cannot be ignored. J Microbiol Immunol Infect. 2020;53(3):444.
- Araújo MPD, Medeiros de Araújo Nunes V, de Albuquerque Costa L, de Souza TA, de Vasconcelos Torres G, Nobre TTX. PLoS One. 2021;16(1):e0245432.
- Sepulveda ER, Stall NM, Sinha SK. A Comparison of COVID-19 Mortality Rates Among Long-Term Care Residents in 12 OECD Countries. J Am Med Dir Assoc. 2020;21(11):1572.
- Lueckmann SL, Hoebel J, Roick J, Markert J, Spallek J, von dem Knesebeck O, et al. Socioeconomic inequalities in primary-care and specialist physician visits: a systematic review. Int J Equity Health. 2021;20(1):1–19.
- Vikum E, Krokstad S, Westin S. Socioeconomic inequalities in health care utilisation in Norway: The population-based HUNT3 survey. Int J Equity Health. 2012;11(1):1–9.
- López RM, Soriano I. Informe Salud y Género 2022: aproximación multidisciplinar a la pandemia por COVID-19. Ministerio de Sanidad, Madrid. 2022.
- Ya'qoub L, Elgendy IY, Pepine CJ. Sex and gender differences in COVID-19: More to be learned! Am Hear J Plus Cardiol Res Pract. 2021;3:100011.
- 31. Ruiz Cantero MT. Health statistics and invisibility by sex and gender during the COVID-19 epidemic. Gac Sanit. 2021;35(1):95–8.
- Martín U, Bacigalupe A, Jiménez Carrillo M. [COVID-19 and gender: certainties and uncertainties in monitoring the pandemic]. Rev Esp Salud Publica. 2021;95:e202104066.
- Akter S. The Gender Gap in COVID-19 Mortality in the United States. Feminist Economics. 2020;27(1–2):30–47.
- Smith J, Davies SE, Grépin KA, Harman S, Herten-Crabb A, Murage A, et al. Reconceptualizing successful pandemic preparedness and response: A feminist perspective. Soc Sci Med. 2022;315:115511.
- Ayittey FK, Dhar BK, Anani G, Chiwero NB. Gendered burdens and impacts of SARS-CoV-2: a review. Health Care Women Int. 2021;41(11–12):1210–25.

- 36. Leon-Larios F, Silva Reus I, Lahoz Pascual I, Quílez Conde JC, Puente Martínez MJ, Gutiérrez Ales J, et al. Women's Access to Sexual and Reproductive Health Services during Confinement Due to the COVID-19 Pandemic in Spain. J Clin Med. 2022;11(14):4074.
- Generales C. Ley Orgánica 3/2007, de 22 de marzo, para la igualdad efectiva de mujeres y hombres. Boe. 2007;71:12611–45.
- WHO. Conferencia Mundial sobre Determinantes Sociales de la Salud Declaración política de Río sobre determinantes sociales de la salud. Available from: https://www.who.int/sdhconference/declaration/Rio_ political_declaration_Spanish.pdf
- Smith J, Davies SE, Feng H, Gan CCR, Grépin KA, Harman S, et al. More than a public health crisis: A feminist political economic analysis of COVID-19. Glob Public Health. 2021;16(8–9):1364–80.
- Wenham C, Davies SE. WHO runs the world–(not) girls: gender neglect during global health emergencies. Int Fem J Polit. 2022;24(3):415–38.

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